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Public Debt, Interest Rates, and Negative Shocks

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This paper studies the broadly measured costs and risks of public debt when it is possible for the government to default. I find that increasing public debt can pose significant macroeconomic and individual welfare risks when rare negative events are possible. These results contrast with recent studies that have argued that the costs of public debt might be low in environments in which interest rates are low for prolonged periods. I use a twoperiod overlapping generations model with aggregate shocks and a lump sum pay-as-you-go government transfer system to study this question.

Friday, December 6h, 2019, 10:30-12:00 Room 126, Extranef building at the University of Lausanne



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GAME OVER: SIMULATING UNSUSTAINABLE FISCAL POLICY

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ABSTRACT

Fiscal sustainability is one of the most pressing policy issues of our time. Yet it remains difficult to quantify. Official debt is plagued with a number of measurement difficulties since its measurement reflects the choice of words, not policies. And forming the fiscal gap–the imbalance in the government's intertemporal budget–requires strong discount rate assumptions. An alternative approach, taken here, is specifying a stochastic general equilibrium model and determining via simulation how long it takes for the economy to reach game over–the point where current policy can no longer be maintained. Our simulations, based on an OLG model calibrated to the U.S. economy, produce an average duration to game over of roughly one century, with a 35 percent chance of reaching the fiscal limit in roughly 30 years. The prospect of man-made economic collapse produces large equity premia, like those observed in the data. Our simulations show that both the fiscal gap and the equity premium rise as the economy gets closer to hitting its fiscal limit, suggesting that the fiscal gap and the equity premium may be good indicators of unsustainable policy.

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1 Introduction

Most developed countries appear to be running unsustainable fiscal policies. In the U.S., federal liabilities (official debt plus the present value of projected non-interest expenditures) exceed federal assets (the present value of projected taxes) by \$211 trillion or 14 times GDP. Closing this fiscal gap requires an immediate and permanent 64 percent hike in all federal taxes.¹ Unlike official debt, the fiscal gap is a label-free and, thus, meaningful measure of fiscal sustainability.² But measuring the fiscal gap raises questions of how to properly discount risky future government purchases and the remaining lifetime net taxes of current and future generations—their generational accounts.

Our approach to assessing sustainability is to simulate a stochastic general equilibrium model and see how long it takes for unsustainable policy to produce game over—the point where the policies can no longer be maintained. Our framework is intentionally simple – a two-period OLG model with first-period labor supply and an aggregate productivity shock. The government redistributes a fixed amount H_t each period from the young to the old. If times become sufficiently bad and the economy reaches game over (i.e., H_t exceeds the earnings of the young), we either let the government take all of the earnings of the young and give them to the old and, thereby, terminate the economy or start redistributing a fixed proportion of earnings from the young to the old.

Our simulations, calibrated to the U.S. economy, produce an average duration to game over of about one century, with a 35 percent chance of reaching the fiscal limit in about 30 years. We also calculate our model's fiscal gap and equity premium. Our model's fiscal gaps are generally small and quite sensitive to the choice of discount rate. But, for any choice of discount factors, the fiscal gaps are much larger when the economy is closer to game over, suggesting that this measure can provide early warning of unsustainable policy.

¹Calculation by authors based on Congressional Budget Office (June 2011) Alternative Fiscal Scenario long-term project of federal cash flows.

²See Kotlikoff and Green (2009).

When post game-over policy terminates the economy, initial period equity premia are about 6 percent—high enough to explain the equity premium puzzle. When gameover is followed by proportional redistribution, equity premiums are initially about 2 percent, but rise dramatically as the economy approaches game over.

When our economy reaches game over, the government is forced to default on its promised payment to the contemporaneous elderly. Thus, this paper contributes to both the literatures on sovereign default³ and fiscally stressed economies.⁴

Our model has no money, so it doesn't include the monetary and fiscal interactions described in Sargent and Wallace (1981) and highlighted in the recent fiscal limits research.⁵ It does include sticky fiscal policy, examined in Alesina and Drazen (1991)as well as Auerbach and Hassett (1992, 2001, 2002, 2007) and Hassett and Metcalf (1999), and regime switching, surveyed in Hamilton (2008).

Section 2 presents the case that game over is followed by policy that kills the economy. Section 3 looks at the switch to policy with either permanently high or moderate intergenerational redistribution. Section 4 concludes.

2 Model with Shut Down

Consider a model with overlapping generations of 2-period-lived agents in which the government redistributes a fixed amount $\bar{H} \geq 0$ from the young to the old each period in which the transfer is feasible. When the transfer is not feasible, the government redistributes all of the available earnings of the young. In so doing, it leaves the economy with no capital in the subsequent period and makes game over economically terminal.

³See Yue (2010), Reinhart and Rogoff (2009), Arellano (2008), Aguiar and Gopinath (2006), Leeper and Walker (2011)

⁴See Auerbach and Kotlikoff (1987), Kotlikoff, Smetters, and Walliser (1998a,b, 2007), İmrohoroğlu, İmrohoroğlu, and Joines (1995, 1999), Huggett and Ventura (1999), Cooley and Soares (1999), De Nardi, İmrohoroğlu, and Sargent (1999), Altig, Auerbach, Kotlikoff, Smetters, and Walliser (2001), Smetters and Walliser (2004), and Nishiyama and Smetters (2007).

⁵See also Cochrane (2011), Leeper and Walker (2011), Davig, Leeper, and Walker (2010, 2011), Davig and Leeper (2011a,b), and Trabandt and Uhlig (2009).

2.1 Household problem

A unit measure of identical agents is born each period. They supply labor only when when young and do so inelastically.

$$l_{1,t} = \bar{l} = 1 \quad \forall t$$

where $l_{1,t}$ is labor supplied by age-1 workers at time t.

Young agents at time t have no wealth and allocate the earnings not extracted by the government between consumption $c_{i,t}$ and saving $k_{i+1,t+1}$ to maximize expected utility. Their problem is

$$\max_{c_{1,t},k_{2,t+1},c_{2,t+1}} u(c_{1,t}) + \beta E_t \left[u(c_{2,t+1}) \right]$$
where $c_{1,t} + k_{2,t+1} \le w_t - H_t$
and $c_{2,t+1} \le (1 + r_{t+1} - \delta)k_{2,t+1} + H_{t+1}$
and $c_{1,t}, c_{2,t+1}, k_{2,t+1} \ge 0$
and where $u(c_{i,t}) = \frac{(c_{i,t})^{1-\gamma} - 1}{1 - \gamma}$

Consumption in the second period of life satisfies

$$c_{2,t+1} = (1 + r_{t+1} - \delta)k_{2,t+1} + H_{t+1} \tag{1}$$

The non-negativity constraint on consumption never binds because each term on the right-hand-side of (1) is weakly positive. Consumption and saving when young, $c_{1,t}$ and $k_{2,t+1}$, are jointly determined by the first-period budget constraint and the Euler equation.

$$c_{1,t} + k_{2,t+1} = w_t - H_t \tag{2}$$

$$u'(c_{1,t}) = \beta E_t \Big[\big(1 + r_{t+1} - \delta \big) u'(c_{2,t+1}) \Big]$$
(3)

From the right-hand-side of (2), the non-negativity constraints on $c_{1,t}$ and $k_{2,t+1}$ bind

when $w_t \leq \overline{H}$. In these cases the government is only able to collect $H_t = w_t$. In so doing, it forces the consumption and saving of the young to zero and terminates the economy.

2.2 Firms' problem

Firms collectively hire labor L_t , at real wage w_t , and rent capital K_t , at real rental rate r_t . Output, Y_t , is produce via the Cobb-Douglas function,

$$Y_t = A_t K_t^{\alpha} L_t^{1-\alpha} \quad \forall t \tag{4}$$

where $A_t = e^{z_t}$ is distributed log normally, and z_t follows an AR(1) process.

$$z_t = \rho z_{t-1} + (1 - \rho)\mu + \varepsilon_t$$
(5)
where $\rho \in [0, 1), \quad \mu \ge 0, \quad \text{and} \quad \varepsilon_t \sim N(0, \sigma^2)$

Profit maximization implies

$$r_t = \alpha e^{z_t} K_t^{\alpha - 1} L_t^{1 - \alpha} \quad \forall t \tag{6}$$

$$w_t = (1 - \alpha)e^{z_t} K_t^{\alpha} L_t^{-\alpha} \quad \forall t \tag{7}$$

2.3 Market clearing

In equilibrium, factor markets clear and national saving equals net investment.

$$L_t = l_1 = \bar{l} = 1 \quad \forall t \tag{8}$$

$$K_t = k_{2,t} \quad \forall t \tag{9}$$

$$Y_t - C_t = K_{t+1} - (1 - \delta)K_t \quad \forall t$$
(10)

where C_t in (10) is aggregate consumption; i.e., $C_t \equiv \sum_{i=1}^2 c_{i,t}$.

2.4 Solution and calibration

A competitive equilibrium for a given \overline{H} is defined as follows.

Definition 1 (Competitive equilibrium). A competitive equilibrium with economic shut down when $w_t < \overline{H}$ is defined as consumption $c_{1,t}$ and $c_{2,t}$ and savings $k_{2,t+1}$ allocations and a real wage w_t and real net interest rate r_t each period such that:

- i. households optimize according to (1), (2) and (3),
- ii. firms optimize according to (6) and (7),
- iii. markets clear according to (8), (9), and (10).

To solve the model, we rewrite (2) as

$$k_{2,t+1} = w_t - H_t - c_{1,t} \tag{11}$$

and use this and the model's other equations to write the Euler equation as

$$u'(c_{1,t}) = \beta E_{z_{t+1}|z_t} \left[\left(1 + \alpha e^{z_{t+1}} \left[(1-\alpha) e^{z_t} k_{2,t}^{\alpha} - \bar{H} - c_{1,t} \right]^{\alpha - 1} - \delta \right] \times \dots \right] u' \left(\left[1 + \alpha e^{z_{t+1}} \left([1-\alpha] e^{z_t} k_{2,t}^{\alpha} - \bar{H} - c_{1,t} \right)^{\alpha - 1} - \delta \right] \left([1-\alpha] e^{z_t} k_{2,t}^{\alpha} - \bar{H} - c_{1,t} \right) + H_{t+1} \right) \right] dt'$$

(12)

where

$$H_t = \min\{w_t, \bar{H}\} = \min\{[1 - \alpha]e^{z_t}k_{2,t}^{\alpha}, \bar{H}\} \quad \forall t$$
(13)

Equations (12) and (13) determine $c_{1,t}$ when $w_t > \overline{H}$. Otherwise, $H_t = w_t$, leaving the young at t with zero consumption and saving $(c_{1,t} = k_{2,t+1} = 0)$.

Given our calibration described in Table 1, which treats one period as 30 years, we solve the above two equations obtaining functions for $c_{1,t}$, $c_{2,t}$, $k_{2,t+1}$, Y_t , w_t , and r_t for any state $(k_{2,t}, z_t)$.⁶

⁶MatLab code for the computation is available upon request.

Parameter	Source to match	Value
β	annual discount factor of 0.96	0.29
γ	coefficient of relative risk aversion between 1.5 and 4.0	2
lpha	capital share of income	0.35
δ	annual capital depreciation of 0.05	0.79
ho	AR(1) persistence of normally distributed shock to match	0.21
	annual persistence of 0.95	
μ	AR(1) long-run average shock level	0
σ	standard deviation of normally distributed shock to match	1.55
	the annual standard deviation of real GDP of 0.49	
\bar{H}	set to be 32% of the median real wage	0.11

Table 1: Calibration of 2-period lived agent OLG model with promised transfer \bar{H}

The Appendix gives a detailed description of the calibration of all parameters.

2.5 Simulation

To explore our model, we ran 3,000 simulations for each of nine combinations of the state variables and \overline{H} . For each of these simulations, we followed the economy through shut down. The nine combinations includes three values of $\overline{H} = \{0.05, 0.11, 0.17\}$ and for three different values of $k_{2,0} = \{0.11, 0.14, 0, 17\}$.⁷ In each simulation we set the initial value of z at its median value μ . Recall that $k_{2,0}$ references the capital held by the old (generation 2) at time zero. Also note that median values refer to the medians taken across all simulations for all periods in which the economy is still functioning.

Table 2 shows the median wage w_{med} , the median capital stock k_{med} , and the size of \bar{H} and $k_{2,0}$ relative to the median wage w_{med} and the median capital stock k_{med} , respectively, for each of the nine combinations of \bar{H} and $k_{2,0}$.

Table 3 provides four statistics on time to economic shutdown, i.e., $w_t \leq \bar{H}$. The middle row of Table 3 corresponding to $\bar{H} = 0.11$ shows that this model economy has a greater than 50 percent chance of shutting down in 60 years (2 periods) under a fiscal transfer system calibrated to be close to that of the United States. Table 3 also indicates what one would expect – that the probability of a near-term shutdown is very sensitive to the size of \bar{H} given the size of the economy's time-zero capital stock

⁷The three values for each roughly correspond to low, middle and high values. That is, $\bar{H} = 0.11$ is the value that is roughly equal to 32 percent of the median wage, and $k_{2,0} = 0.14$ is roughly equal to the median capital stock across simulations.

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	k _{2,0} =	$k_{2,0} = 0.11$		= 0.14	$k_{2,0} = 0.17$	
	w_{med}	k_{med}	w_{med}	k_{med}	w_{med}	k_{med}
	\bar{H}/w_{med}	$k_{2,0}/k_{med}$	\bar{H}/w_{med}	$k_{2,0}/k_{med}$	\bar{H}/w_{med}	$k_{2,0}/k_{med}$
$\bar{H} = 0.05$	0.3030	0.0992	0.3026	0.0996	0.3008	0.0991
II = 0.05	0.1650	1.1093	0.1652	1.4062	0.1662	1.7148
$\bar{H} = 0.11$	0.3445	0.1344	0.3433	0.1358	0.3474	0.1365
II = 0.11	0.3193	0.8187	0.3204	1.0311	0.3166	1.2457
$\bar{H} = 0.17$	0.2562	0.1043	0.2709	0.1090	0.2825	0.1134
11 = 0.17	0.6635	1.0550	0.6275	1.2846	0.6018	1.4988

Table 2: Initial values relative to median values

 w_{med} is the median wage and k_{med} is the median capital stock across all 3,000 simulations before economic shut down.

and, thus, initial wage.

		$k_{2,0} = 0.11$		$k_{2,0} = 0.14$		$k_{2,0} = 0.17$	
		Periods	CDF	Periods	CDF	Periods	CDF
	min	1	0.1620	1	0.1543	1	0.1477
$\bar{U} = 0.05$	med	4	0.5370	4	0.5320	4	0.5283
H = 0.05	mean	5.95	0.6704	6.00	0.6703	6.04	0.6694
	max	45	1.0000	45	1.0000	45	1.0000
	min	1	0.3623	1	0.3480	1	0.3357
$\bar{U} = 0.11$	med	2	0.5653	2	0.5543	2	0.5433
II = 0.11	mean	3.29	0.7060	3.35	0.7029	3.41	0.7022
	max	24	1.0000	24	1.0000	25	1.0000
	min	1	0.5203	1	0.4987	1	0.4807
$\bar{H} = 0.17$	med	1	0.5203	2	0.6833	2	0.6707
	mean	2.42	0.7373	2.48	0.7336	2.54	0.7295
	\max	18	1.0000	18	1.0000	18	1.0000

Table 3: Periods to shut down simulation statistics

The "min", "med", "mean", and "max" rows in the "Periods" column represent the minimum, median, mean, and maximum number of periods, respectively, in which the simulated time series hit the economic shut down. The "CDF" column represents the percent of simulations that shut down in t periods or less, where t is the value in the "Periods" column. For the CDF value of the "mean" row, we used linear interpolation.

2.6 Fiscal gap and equity premium

Because actual receipts extracted from young workers are not always equal to the promised payment $H_t \leq \overline{H}$, we define the fiscal gap as the difference between the present value of all promised payments to current and future older generations and

the present value of current receipts from current workers plus all future receipts obtained, on average in each future period, from future workers. We express this difference as a percent of the present value of all current and future output realized, on average.

fiscal gap_t =
$$x_t \equiv \frac{NPV(\bar{H}) - NPV(H_t)}{NPV(Y_t)}$$
 (14)

This measure does not suffer from the economics labeling problem.

Define the discount factor in s periods from the current period as d_{t+s} , and write the net present values in the measure of the fiscal gap from (14) in terms of the discount factors and expected streams of transfers and income.

$$x_{t} = \frac{\sum_{s=0}^{\infty} d_{t+s} \bar{H} - \sum_{s=0}^{\infty} d_{t+s} E\left[H_{s}\right]}{\sum_{s=0}^{\infty} d_{t+s} E\left[Y_{s}\right]}$$
(15)

We present four measures of the fiscal gap using four sequences of discount factors d_{t+s} —two from our model and two from the literature. The first measure of the fiscal gap (fgap1) uses the prices of sure-return bonds that mature s periods from the current period t as the discount factors. Define $p_{t,j}$ as the price of an asset $B_{t,j}$ with a sure-return payment of one unit j periods in the future. If these assets can be bought and sold each period, then a household could purchase an asset that pays off after the household is dead and sell it before they die. Because each of these assets must be held in zero net supply, they do not change the equilibrium policy functions described in Section 2.4. The equations characterizing the prices $p_{t,j}$ for all t and j are:⁸

$$p_{t,j} = \begin{cases} 1 & \text{if } j = 0 \\ \beta \frac{E_t[u'(c_{2,t+1})p_{t+1,j-1}]}{u'(c_{1,t})} & \text{if } j \ge 1 \end{cases} \quad \forall t$$
(16)

With the starting value of the sure-return price $p_{t,0}$ pinned down, the prices of the assets that mature in future periods can be calculated recursively using equation (16).

Table 4 shows the calculated sure-return prices at each maturity—which we use as our discount factors—and their corresponding net discount rates shown on an annual

⁸We derive equation (16), as well as some other assets of interest, in detail in the Technical Appendix.

basis. The first column in each cell displays the prices of the different maturity s of sure return bond $p_{t,t+s}$ computed using recursive equation (16). The second column in each cell represents the annualized version of the net return $r_{t,t+s}$ APR or net interest rate.⁹

$$r_{t,t+s} = \left(\frac{1}{p_{t,t+s}}\right)^{\frac{1}{s_{30}}} - 1 \quad \text{for} \quad s \ge 1$$
 (17)

		$k_{2,0} =$	0.11	$k_{2,0} =$	0.14	$k_{2,0} =$	0.17
			$r_{t,t+s}$		$r_{t,t+s}$		$r_{t,t+s}$
_	s	$p_{t,t+s}$	APR	$p_{t,t+s}$	APR	$p_{t,t+s}$	APR
	0	1	0	1	0	1	0
	1	1.5556	-0.0146	1.5897	-0.0153	1.6190	-0.0159
Ū 0.05	2	0.3115	0.0196	0.3466	0.0178	0.3782	0.0163
	3	0.0385	0.0369	0.0441	0.0353	0.0493	0.0340
II = 0.03	4	0.0088	0.0403	0.0096	0.0395	0.0099	0.0392
	5	0.0049	0.0360	0.0063	0.0344	0.0063	0.0344
	6	0.0014	0.0372	0.0025	0.0338	0.0024	0.0342
	0	1	0	1	0	1	0
	1	1.6771	-0.0171	1.7186	-0.0179	1.7673	-0.0188
	2	0.1543	0.0316	0.1793	0.0291	0.2137	0.0261
$\overline{U} = 0.11$	3	0.0074	0.0560	0.0092	0.0535	0.0118	0.0506
$\Pi = 0.11$	4	0.0072	0.0420	0.0077	0.0414	0.0085	0.0405
	5	0.0029	0.0397	0.0032	0.0390	0.0038	0.0379
	6	4.3×10^{-4}	0.0440	5.0×10^{-4}	0.0431	5.9×10^{-4}	0.0421
	0	1	0	1	0	1	0
	1	1.5848	-0.0152	1.6811	-0.0172	1.7308	-0.0181
	2	0.0092	0.0812	0.0156	0.0718	0.0359	0.0570
\overline{U} 0.17	3	0.0010	0.0794	0.0031	0.0663	0.0038	0.0639
$\Pi = 0.17$	4	9.0×10^{-5}	0.0808	0.0046	0.0459	0.0049	0.0453
	5	1.3×10^{-5}	0.0780	0.0010	0.0470	0.0011	0.0463
	6	1.7×10^{-5}	0.0630	5.6×10^{-5}	0.0558	6.1×10^{-5}	0.0554

Table 4: Term structure of prices and interest rates

The first column in each cell is the price of the sure-return bond $p_{t,t+s}$ at different maturities s as characterized by equation (16). The second column in each cell is the net interest rate $r_{t,t+s}$ APR implied by the sure-return rate and given in annual percentage rate terms according to equation (17). Full descriptions of the term structure of prices and interest rates for all calibrations and for up to s = 12 is provided in the Technical Appendix.

The second fiscal gap measure (fgap 2) employs a constant discount rate, namely

⁹The return or yield of a sure-return bond should increase with its maturity in an economy that never shuts down. However, the increasing probability of the economy shutting down in each future period counteracts the increasing value of the sure return in the future. This is why the interest rates in the second column of each cell in Table 4 seem to go toward an asymptote in the limit.

the current-period risky return on capital R_t . For example, the risky return on capital in period t is $R_t = 1.4971$ in the middle cell in which $\bar{H} = 0.11$ and $k_{2,0} = 0.14$. So the discount factors are $d_{t+s} = (1.4971)^{-s}$. Our third fiscal gap measure (fgap 3) uses a constant discount rate taken from International Monetary Fund (2009, Table 6.4). This study uses an annual discount factor of the growth rate in real GDP plus 1 percent to calculate the net present value of aging-related expeditures. This averages out across G-20 countries to be a discount rate of around 4 percent and for the U.S. is about 3.8 percent ($R_t \approx 3.1$). So the discount rates for fgap3 are $d_{t+s} = (3.05)^{-s}$. For the last measure of the fiscal gap (fgap4), we use the constant discount rate from Gohkhale and Smetters (2007) who use an annual discount rate of 3.65 percent for their discount factors in their NPV calculation. This is equivalent to a 30-year gross discount rate of $R_t \approx 2.9$. So the discount rates for fgap4 are $d_{t+s} = (2.93)^{-s}$. The expectations for H_t and Y_t are simply the average values from the 3,000 simulations described in Section 2.5.

Table 5 presents fiscal gaps for the nine different combinations of promised transfers \bar{H} and initial capital stock $k_{2,0}$ as a percent of the net present value of output. By way of comparison, we note that the U.S. fiscal gap is currently 12 percent of the present value of projected GDP. The figures in table 5 are generally much smaller. Importantly, though, given the initial capital stock, higher values of \bar{H} are associated not just with much quicker time to shut down, but also substantially larger fiscal gaps regardless of the discount rates used.

Next we use the difference in the expected risky return on capital $E[R_{t+1}]$ and the riskless return on the one-period safe bond $R_{t,t+1}$ to calculate an equity premium. A large literature attempts to explain why the observed equity premium is so large.¹⁰ Most recently, Barro (2009) has shown that incorporating rare disasters into an economic model produces realistic risk premia and risk free rates. Our model features disaster in the form of economic shutdown, and it too (see Table 6) model produces realistic equity premia, ranging from 4.7 percent to as 7.3 percent, for a

¹⁰See Shiller (1982), Mehra and Prescott (1985), Kocherlakota (1996), Campbell (2000), and Cochrane (2005, Ch. 21) for surveys of the equity premium puzzle.

	$k_{2,0} = 0.11$		$k_{2,0} =$	= 0.14	$k_{2,0} = 0.17$	
	fgap 1	fgap 2	fgap 1	fgap 2	fgap 1	fgap 2
	fgap 3	fgap 4	fgap 3	fgap 4	fgap 3	fgap 4
$\bar{H} = 0.05$	0.0037	0.0078	0.0034	0.0096	0.0033	0.0118
11 - 0.05	0.0033	0.0035	0.0030	0.0032	0.0028	0.0029
$\bar{H} = 0.11$	0.0192	0.0373	0.0175	0.0427	0.0164	0.555
II = 0.11	0.0168	0.0176	0.0152	0.0159	0.0140	0.0147
$\bar{H} = 0.17$	0.0474	0.0876	0.0421	0.1041	0.0385	0.1171
11 = 0.17	0.0408	0.0426	0.0361	0.0378	0.0328	0.0344

Table 5: Measures of the fiscal gap as percent of NPV(GDP)

Fiscal gap 1 uses the gross sure return rates $R_{t,t+s}$ from Table 4 as the discount rates for NPV calculation. Fiscal gap 2 uses the current period gross return on capital R_t from the model as the constant discount rate. Fiscal gap 3 uses the International Monetary Fund (2009) method of an annual discount rate equal to 1 plus the average percent change in GDP plus 0.01 (\approx 2.05). And fiscal gap 4 uses the Gohkhale and Smetters (2007) method of an annual discount rate equal to 1 plus 0.0365 (\approx 1.93).

moderate-sized coefficient of relative risk aversion of $\gamma = 2$.

Table 6 present Sharpe ratios as well as all of the components of the equity premium. For the expected risky return $E[R_{t+1}]$, the one-period sure return $R_{t,t+1}$, and the equity premium (the difference between the two), we report results for both one period from the model (30 years) as well as the annualized (one-year) version. Our Sharpe ratios between 0.32 and 0.33 are in line with common estimates from the data.

Because the equity premium and the Sharpe ratio fluctuate from period-to-period, we report in Table 7 the average equity premium and Sharpe ratio across simulations in the period immediately before the economic shutdown.

Table 8 compares three measures of the fiscal gap in the initial period to the fiscal gap in the period immediately before shutdown.¹¹ As with the equity premium and Sharpe ratio, the fiscal gap increases on average in the period immediately before shutdown relative to the initial period. In our baseline case of $\bar{H} = 0.11$ and $k_{2,0} = 0.14$, the fiscal gap nearly doubles in the period before shutdown. Tables 7 and 8 provide evidence that both the fiscal gap and the equity premium are good leading indicators of how close an economy is to its fiscal limit.

¹¹We exclude the calculation of measure of the fiscal gap that uses the current period marginal product of capital as the discount rate (fgap2) because the discount rate is often negative in the period immediately before shutdown. We exclude fgap3 because it is similar to fgap4.

		$k_{2,0} =$	- 0.11	$k_{2,0} =$	0.14	$k_{2,0} = 0.17$	
		30-year	annual	30-year	annual	30-year	annual
	$E[R_{t+1}]$	8.2070	1.0361	7.5150	1.0334	7.0113	1.0313
	$\sigma(R_{t+1})$	23.3433	n.a.	21.3222	n.a.	19.8511	n.a.
	$R_{t,t+1}$	0.6428	0.9854	0.6291	0.9847	0.6177	0.9841
$\bar{H} = 0.05$	Equity premium $E[R_{t+1}] - R_{t,t+1}$	7.5641	0.0507	6.8859	0.0487	6.3936	0.0473
	Sharpe ratio $\frac{E[R_{t+1}] - R_{t,t+1}}{\sigma(R_{t+1})}$	0.3240	n.a.	0.3229	n.a.	0.3221	n.a.
	$E[R_{t+1}]$	11.3042	1.0459	10.0769	1.0423	9.2241	1.0396
	$\sigma(R_{t+1})$	32.3859	n.a.	28.8049	n.a.	26.3140	n.a.
	$R_{t,t+1}$	0.5963	0.9829	0.5819	0.9821	0.5658	0.9812
$\bar{H} = 0.11$	Equity premium $E[R_{t+1}] - R_{t,t+1}$	10.7080	0.0630	9.4950	0.0602	8.6582	0.0584
	Sharpe ratio $\frac{E[R_{t+1}] - R_{t,t+1}}{\sigma(R_{t+1})}$	0.3306	n.a.	0.3296	n.a.	0.3290	n.a.
	$E[R_{t+1}]$	16.2082	1.0574	13.7520	1.0521	12.1889	1.0483
	$\sigma(R_{t+1})$	46.7126	n.a.	39.5389	n.a.	34.9735	n.a.
	$R_{t,t+1}$	0.6310	0.9848	0.5948	0.9828	0.5778	0.9819
$\bar{H} = 0.17$	Equity premium $E[R_{t+1}] - R_{t,t+1}$	15.5772	0.0727	13.1572	0.0693	11.6112	0.0664
	Sharpe ratio $\frac{E[R_{t+1}] - R_{t,t+1}}{\sigma(R_{t+1})}$	0.3335	n.a.	0.3328	n.a.	0.3320	n.a.

Table 6: Components of the equity premium in period 1

The gross risky one-period return on capital is $R_{t+1} = 1 + r_{t+1} - \delta$. The annualized gross risky one-period return is $(R_{t+1})^{1/30}$. The expected value and standard deviation of the gross risky one-period return R_{t+1} are calculated as the average and standard deviation, respectively, across simulations. The annual equity premium is the expected value of the annualized risky return in the next period minus the annualized return on the one-period riskless bond.

		$k_{2,0} =$	= 0.11	$k_{2,0} =$	= 0.14	k _{2,0} =	= 0.17
		Eq.	Sharpe	Eq.	Sharpe	Eq.	Sharpe
		prem.	ratio	prem.	ratio	prem.	ratio
	period 1	0.0507	0.3240	0.0487	0.3229	0.0473	0.3221
$\bar{H} = 0.05$	before shutdown	0.0710	0.3356	0.0707	0.3337	0.0706	0.3370
II = 0.05	percent bigger	0.6617	0.5410	0.6843	0.5570	0.6960	0.5690
	percent smaller	0.1763	0.2970	0.1613	0.2887	0.1563	0.2833
	period 1	0.0630	0.3306	0.0602	0.3296	0.0584	0.3290
$\bar{H} = 0.11$	before shutdown	0.0679	0.3339	0.0667	0.3333	0.0664	0.3343
II = 0.11	percent bigger	0.3740	0.3760	0.4023	0.3970	0.4227	0.4153
	percent smaller	0.2637	0.2617	0.2497	0.2550	0.2417	0.2490
	period 1	0.0727	0.3335	0.0693	0.3328	0.0664	0.3320
\bar{H} 0.17	before shutdown	0.0709	0.3353	0.0686	0.3354	0.0673	0.3348
II = 0.11	percent bigger	0.2027	0.2740	0.2253	0.2937	0.2543	0.3070
	percent smaller	0.2770	0.2057	0.2760	0.2077	0.2650	0.2123

 Table 7: Equity premium and Sharpe ratio in period immediately before shutdown

The "period 1" row represents the equity premium and Sharpe ratio in the initial period for each specification. The "before shutdown" row represents the average equity premium and Sharpe ratio across simulations in the period immediately before shutdown for each specification. The "percent bigger" and "percent smaller" rows tell how many of the simulated ending values of the equity premium and Sharpe ratio were bigger than or less than, respectively, their initial period values. These percentages do not sum to one because the equity premium and Sharpe ratio do not change in the cases in which the economy shuts down in the second period.

		$k_{2,0} =$	= 0.11	$k_{2,0} =$	= 0.14	$k_{2,0} =$	= 0.17
		fgap1	fgap4	fgap1	fgap4	fgap1	fgap4
	period 1	0.0037	0.0035	0.0034	0.0032	0.0033	0.0029
$\bar{H} = 0.05$	before shutdown	0.0183	0.0187	0.0184	0.0187	0.0185	0.0191
II = 0.05	percent bigger	0.8940	0.8260	0.9100	0.8340	0.9220	0.8440
	percent smaller	0.1060	0.1740	0.0900	0.1660	0.0780	0.1560
	period 1	0.0192	0.0176	0.0175	0.0159	0.0164	0.0147
$\bar{H} = 0.11$	before shutdown	0.0339	0.0291	0.0337	0.0294	0.0356	0.0307
II = 0.11	percent bigger	0.7200	0.6940	0.7480	0.7060	0.7600	0.7160
	percent smaller	0.2800	0.3060	0.2520	0.2940	0.2400	0.2840
	period 1	0.0474	0.0426	0.0421	0.0378	0.0385	0.0344
\overline{U} 0.17	before shutdown	0.0508	0.0447	0.0481	0.0429	0.0495	0.0414
11 - 0.17	percent bigger	0.7180	0.6820	0.7340	0.6760	0.7500	0.6824
	percent smaller	0.2820	0.3180	0.2660	0.3240	0.2500	0.3180

Table 8: Fiscal gaps in period immediately before shutdown

The "period 1" row represents the fiscal gap in the initial period for each specification. The "before shutdown" row represents the average fiscal gap across simulations in the period immediately before shutdown for each specification. The "percent bigger" and "percent smaller" rows tell how many of the simulated ending values of the fiscal gap were bigger than or less than, respectively, their initial period values. Fiscal gap 1 uses the gross sure return rates $R_{t,t+s}$ similar to Table 4 as the discount rates for NPV calculation, and fiscal gap 4 uses the Gohkhale and Smetters (2007) method of an annual discount rate equal to 1 plus 0.0365 (≈ 1.93).

3 Model with Regime Change

We now assume that when the government defaults on its promised transfer $w_t \leq \bar{H}$, the regime switches permanently to one in which the transfer is simply τ percent of the wage each period $H_t = \tau w_t$. We solve the model for $\tau = 0.8$ and $\tau = 0.3$.

3.1 Regime change to 80-percent wage tax

Figure 1 illustrates the rule for the transfer H_t under regime 1 in which the transfer is \overline{H} unless wages w_t are less than \overline{H} and under regime 2 in which the transfer is permanently switched to the proportional transfer system $H_t = 0.8w_t$.

3.1.1 Household problem, firm problem, and market clearing

The characterization of the household problem remains the same as in equations (1), (2), and (3) from Section 2.1. The only difference is in the definition of H_t in those equations. With the new regime switching assumption, the transfer each period from

Figure 1: Transfer program H_t under regime 1 and regime 2: 80 percent wage tax



the young to the old H_t is defined as follows.

$$H_t = \begin{cases} \bar{H} & \text{if } w_s > \bar{H} & \text{for all} \quad s \le t \\ 0.8w_t & \text{if } w_s \le \bar{H} & \text{for any} \quad s \le t \end{cases}$$
(18)

The change is reflected in the expectations of the young of consumption when old $c_{2,t+1}$ in the savings decision (3).

The firm's problem and the characterization of output, aggregate productivity shock, and optimal net real return on capital and real wage are the same as equations (4) through (7) in Section 2.2. The market clearing conditions that must hold in each period are the same as (8), (9), and (10) from Section 2.3

3.1.2 Solution and calibration

The competitive equilibrium with a transfer program regime switch is characterized in the same way as Definition 1 with economic shut down except that the transfer each period is characterized by equation (18). For the current young, this regime switch decreases the expected value of next period's transfer $H_{t+1} - 0.8w_t + 1$ instead of $w_t + 1$. Thus, the current period young will save more and bring more savings $k_{2,t+1}$ into old age than did the young in Section 2. Once the regime has permanently switched to the high-rate proportional transfer program of $H_t = 0.8w_t$, allocations each period are determined by the following two equations,

$$c_{2,t} = (1 + \alpha e^{z_t} k_{2,t}^{\alpha - 1} - \delta) k_{2,t} + 0.8(1 - \alpha) e^{z_t} k_{2,t}^{\alpha}$$
(19)

$$u'(c_{1,t}) = \beta E_{z_{t+1}|z_t} \left[\left(1 + \alpha e^{z_{t+1}} k_{2,t+1}^{\alpha - 1} - \delta \right) \times \dots \right] u'\left(\left[1 + \alpha e^{z_{t+1}} k_{2,t+1}^{\alpha - 1} - \delta \right] k_{2,t+1} + 0.8(1 - \alpha) e^{z_{t+1}} k_{2,t+1}^{\alpha} \right) \right]$$
(20)

where,

$$k_{2,t+1} = 0.2(1-\alpha)e^{z_t}k_{2,t}^{\alpha} - c_{1,t}$$
(21)

and in which we have substituted in the expressions for r_t and w_t from (6) and (7), respectively, and $H_t = 0.8w_t$.

We calibrate parameters as in Table 1 for the economic shut down model with the exception of \bar{H} . We again calibrate \bar{H} to be 32 percent of the median wage. However, we calculate the median wage from the time periods in the simulations before the regime switches (regime 1). Because the economy never shuts down, it is less risky in the long run. But the economy is actually more risky to the current period young in that the expected value of their transfer in the next period is decreased by the potential regime switch. Higher precautionary saving induces a higher median wage and a higher promised transfer $\bar{H} = 0.09$ in order to equal 32 percent of the regime 1 median wage.

3.1.3 Simulation

We again simulate the regime switching model 3,000 times with various combinations of values for the promised transfer $\bar{H} \in \{0.09, 0.11\}$ and the initial capital stock $k_{2,0} \in \{0.0875, 0.14\}$. As shown in Table 9, our calibrated values of $\bar{H} = 0.09$ and $k_{2,0} = 0.0875$ correspond to 32 percent of the median real wage in regime 1 and the median capital stock in regime 1, respectively. In each simulation we again use an initial value of the productivity shock of its median value $z_0 = \mu$.

The upper left cell of Table 9 is analogous to the middle cell of Table 2 in that \overline{H}

	$k_{2,0} = 0.0875$		$k_{2,0}$ =	= 0.14	
	w_{med}	k_{med}	w_{med}	k_{med}	
	\bar{H}/w_{med}	$k_{2,0}/k_{med}$	\bar{H}/w_{med}	$k_{2,0}/k_{med}$	
\overline{U} 0.00	0.2827	0.0878	0.2883	0.0895	
II = 0.09	0.3184	0.9967	0.3121	1.5642	
$\bar{H} = 0.11$	0.2944	0.0886	0.3021	0.0899	
	0.3736	0.9873	0.3641	1.5567	

Table 9: Initial values relative to median values from regime 1: 80-percent tax

 w_{med} is the median wage and k_{med} is the median capital stock across all 3,000 simulations before the regime switch (in regime 1).

is calibrated to be 32 percent of the regime 1 real wage and $k_{2,0}$ to equal the regime 1 median capital stock. However, the lower right cell of Table 9 has the same \bar{H} and $k_{2,0}$ as the middle cell of Table 2. Notice that the median capital stock is higher in the regime switching economy ($k_{med} = 0.1.5567$ for $\bar{H} = 0.11$ and $k_{2,0} = 0.14$ in regime switching economy as compared to $k_{med} = 0.1.0311$ in the shutdown economy with the same \bar{H} and $k_{2,0}$). This is because young households have an increased risk in the second period of life under the possibility of a regime switch because their transfer will be lower in the case of a default on \bar{H} .

Table 10 presents time to game over for this policy. Notice that the distribution of time until regime switch across simulations from the upper left cell of Table 10 is very similar to the middle cell in Table 3 from the shut down economy. Higher precautionary savings extends the time until a regime switch, but increased promised transfers reduce that time.

3.1.4 Fiscal gap and equity premium

For the model with regime switching to an 80-percent wage tax, we define the fiscal gap in the same way as in equation (14) from Section 2.6. The discount factors used to calculate the net present values in the fiscal gap measures from the regime switching model are calculated in the same way as described in Section 2.6. Table 11 shows the calculated sure-return prices and their corresponding annualized discount rates for this regime switching economy. Each cell represents the computed prices and interest

		$k_{2,0} = 0.0875$		$k_{2,0} =$	0.14
		Periods	CDF	Periods	CDF
	\min	1	0.3677	1	0.3340
$\bar{H} = 0.09$	med	2	0.5727	2	0.5470
	mean	3.25	0.7124	3.40	0.7066
	\max	24	1.0000	25	1.0000
	\min	1	0.4517	1	0.4060
$\bar{H} = 0.11$	med	2	0.6430	2	0.6127
	mean	2.78	0.7314	2.94	0.7244
	\max	24	1.0000	24	1.0000

Table 10: Periods to regime switch simulation statistics: 80-percent tax

The "min", "med", "mean", and "max" rows in the "Periods" column represent the minimum, median, mean, and maximum number of periods, respectively, in which the simulated time series hit the regime switch condition. The "CDF" column represents the percent of simulations that switch regimes in t periods or less, where t is the value in the "Periods" column. For the CDF value of the "mean" row, we used linear interpolation.

rates that correspond to a particular promised transfer value \bar{H} and initial capital stock $k_{2,0}$.

Table 12 shows our four measures of the fiscal gap as a percent of the net present value of GDP for each of our four combinations of \bar{H} and $k_{2,0}$. Some of the fiscal gap measures are negative. This occurs because some of the discount factors decay more slowly than others (fgap 1 is the slowest) and because expected receipts are higher after the regime switch. Indeed, they can even end up higher than than \bar{H} . Even though the impulse response of w_t decays to a lower level after the regime switch, the expected H_t can be high because of the high variance in productivity shocks. A median value would be lower. We therefore can get negative fiscal gap measures, even though \bar{H} is big enough to trigger a regime switch in relatively few periods. Table 12 computes fiscal gaps as a percent of the present value of output as in equation (14) for the four combinations of values for the promised transfer \bar{H} and the initial capital stock $k_{2,0}$.

Note also in Table 12 that the fiscal gap measure fgap1 becomes even more negative as \bar{H} increases. This is caused by the higher \bar{H} shortening the periods until the regime switch or higher H_t values. In other words, the positive effect on the fiscal gap from

		$k_{2,0} =$	0.0875	$k_{2,0}$ =	= 0.14
			$r_{t,t+s}$		$r_{t,t+s}$
	s	$p_{t,t+s}$	APR	$p_{t,t+s}$	APR
	0	1	0	1	0
	1	0.3269	0.0380	0.4645	0.0259
	2	1.1607	-0.0025	2.5547	-0.0155
$\bar{H} = 0.09$	3	0.3534	0.0116	0.4138	0.0099
	4	0.6753	0.0033	1.2121	-0.0016
	5	0.4117	0.0059	0.2982	0.0081
	6	0.1304	0.0114	0.4420	0.0045
	0	1	0	1	0
	1	0.2328	0.0498	0.3227	0.0384
	2	1.3063	-0.0044	1.5334	-0.0071
$\bar{H} = 0.11$	3	2.5521	-0.0104	1.5811	-0.0051
	4	0.2606	0.0113	0.8424	0.0014
	5	1.7532	-0.0037	1.8832	-0.0042
	6	0.3762	0.0054	0.4895	0.0040

Table 11: Term structure of prices and interest rates in regime switching economy: 80-percent tax

The first column in each cell is the price of the sure-return bond $p_{t,t+s}$ at different maturities s as characterized by equation (16). The second column in each cell is the net interest rate $r_{t,t+s}$ APR implied by the sure-return rate and given in annual percentage rate terms according to equation (17). Full descriptions of the term structure of prices and interest rates for all calibrations and for up to s = 12 is provided in the Technical Appendix.

Table 12: Measures of the fiscal gap with regime switching as percent of NPV(GDP): 80percent tax

	$k_{2,0} =$	0.0875	$k_{2,0} = 0.14$		
	fgap 1	fgap 2	fgap 1	fgap 2	
	fgap 3	fgap 4	fgap 3	fgap 4	
$\bar{H} = 0.00$	-0.0519	0.0003	-0.0343	-0.0157	
11 - 0.09	0.0067	0.0066	0.0052	0.0051	
$\bar{H} = 0.11$	-0.0861	0.0057	-0.0749	-0.0075	
11 - 0.11	0.0130	0.0129	0.0103	0.0102	

Fiscal gap 1 uses the sure return rates $R_{t,t+s}$ from Table 4 to form the discount factors used in its present value calculations. Fiscal gap 2 uses the current period gross return on capital R_t from the model as the constant discount rate. Fiscal gap 3 uses the International Monetary Fund (2009) method of an annual discount rate equal to 1 plus the average percent change in GDP plus 0.01 (\approx 2.05). And fiscal gap 4 follows Gohkhale and Smetters (2007) in forming the discount factors using an annual discount rate equal to 1 plus 0.0365 (\approx 1.93).

a higher \bar{H} in the pre-switch periods is dominated by the negative effect on the fiscal gap from more periods of high regime 2 H_t . For the other measures of the fiscal gap, the second effect dominates so the fiscal gap increases with the size of the promised transfer \bar{H} .

Finally, we calculate the equity premium and Sharpe ratio for this regime-switching model using the difference in the expected risky return on capital one period from now $E[R_{t+1}]$ and the riskless return on the sure-return bond maturing one period from now $R_{t,t+1}$. In reference to the Barro (2009) result, our model with regime switching delivers equity premia that are significantly lower than the riskier model with shut down from Section 2.6 and do not match as closely observed equity premia and Sharpe ratios. As shown in Table 13, our regime switching model produces equity premia around 2 percent and Sharpe ratios around 0.28.

The interesting equity premium story in the model with the 80-percent regime switch is what happens to the equity premium as the economy approaches game one with respect to its initial policy. Table 14 reports the average equity premium and Sharpe ratio across simulations in the period immediately before the regime switch as

		$k_{2,0} = 0.0875$		$k_{2,0} = 0.14$	
		30-year	annual	30-year	annual
	$E[R_{t+1}]$	17.1319	1.0592	12.9708	1.0503
	$\sigma(R_{t+1})$	49.4105	n.a.	37.2570	n.a.
	$R_{t,t+1}$	3.0589	1.0380	2.1526	1.0259
$\bar{H} = 0.09$	Equity premium $E[R_{t+1}] - R_{t,t+1}$	14.0731	0.0213	10.8182	0.0244
	Sharpe ratio $\frac{E[R_{t+1}] - R_{t,t+1}}{\sigma(R_{t+1})}$	0.2848	n.a.	0.2904	n.a.
	$E[R_{t+1}]$	22.1773	1.0678	16.0801	1.0572
	$\sigma(R_{t+1})$	64.1466	n.a.	46.3385	n.a.
	$R_{t,t+1}$	4.2960	1.0498	3.0985	1.0384
$\bar{H} = 0.11$	Equity premium $E[R_{t+1}] - R_{t,t+1}$	17.8813	0.0180	12.9816	0.0188
	Sharpe ratio $\frac{E[R_{t+1}] - R_{t,t+1}}{\sigma(R_{t+1})}$	0.2788	n.a.	0.2801	n.a.

Table 13: Components of the equity premium with
regime switching: 80-percent tax

The gross risky one-period return on capital is $R_{t+1} = 1 + r_{t+1} - \delta$. The annualized gross risky one-period return is $(R_{t+1})^{1/30}$. The expected value and standard deviation of the gross risky one-period return R_{t+1} are calculated as the average and standard deviation, respectively, across simulations. The annual equity premium is the expected value of the annualized risky return in the next period minus the annualized return on the one-period riskless bond. compared to their respective values in the first period. The average equity premium and Sharpe ratio increase significantly from the initial period to the period right before the regime switch in each case.

		$k_{2,0} = 0.0875$		$k_{2,0} = 0.14$	
		Eq.	Sharpe	Eq.	Sharpe
		prem.	ratio	prem.	ratio
	period 1	0.0213	0.2848	0.0244	0.2904
$\bar{H} = 0.00$	before shutdown	0.0737	0.3231	0.0773	0.3272
II = 0.09	percent bigger	0.6287	0.5353	0.6600	0.5523
	percent smaller	0.0037	0.0970	0.0060	0.1137
$\bar{H} = 0.11$	period 1	0.0180	0.2788	0.0188	0.2801
	before shutdown	0.0637	0.3152	0.0675	0.3201
	percent bigger	0.5457	0.4770	0.5910	0.5180
	percent smaller	0.0027	0.0713	0.0030	0.0760

Table 14: Equity premium and Sharpe ratio in period immediatelybefore regime switch: 80-percent tax

The "period 1" row represents the equity premium and Sharpe ratio in the initial period for each specification. The "before shutdown" row represents the average equity premium and Sharpe ratio across simulations in the period immediately before shutdown for each specification. The "percent bigger" and "percent smaller" rows tell how many of the simulated ending values of the equity premium and Sharpe ratio were bigger than or less than, respectively, their initial period values. These percentages do not sum to one because the equity premium and Sharpe ratio do not change in the cases in which the economy shuts down in the second period.

3.2 Regime change to 30-percent wage tax

In this section, we show the effects of a less severe proportional wage tax of 30 percent $H_t = 0.3w_t$ in the case of a regime switch.

3.2.1 Simulation

As shown in Table 15, our calibrated values of $\bar{H} = 0.09$ and $k_{2,0} = 0.0875$ again correspond to about 32 percent of the median real wage in regime 1 and close to the median capital stock in regime 1, respectively. Note that none of these regime 1 values change much from Table 9 even though regime 2 entails switching to a very different policy when current policy fails. In each simulation we use an initial value of the productivity shock of its median value $z_0 = \mu$.

	$k_{2,0} = 0.0875$		$k_{2,0}$ =	= 0.14	
	w_{med} k_{med}		w_{med}	k_{med}	
	\bar{H}/w_{med}	$k_{2,0}/k_{med}$	\bar{H}/w_{med}	$k_{2,0}/k_{med}$	
$\bar{H} = 0.00$	0.2828	0.0864	0.2880	0.0885	
II = 0.09	0.3183	1.0130	0.3125	1.5819	
$\bar{U} = 0.11$	0.2963	0.0868	0.3051	0.0877	
H = 0.11	0.3712	1.0082	0.3605	1.5970	

Table 15: Initial values relative to median values from regime 1: 30-percent tax

 w_{med} is the median wage and k_{med} is the median capital stock across all 3,000 simulations before the regime switch (in regime 1).

The upper left cell of Table 15 is analogous to the middle cell of Table 2 in that \bar{H} is calibrated to be 32 percent of the regime 1 real wage and $k_{2,0}$ to equal the regime 1 median capital stock. However, the lower right cell of Table 15 has the same \bar{H} and $k_{2,0}$ as the middle cell of Table 2. Notice that the median capital stock is higher in the regime-switching economy ($k_{med} = 0.1.5970$ for $\bar{H} = 0.11$ and $k_{2,0} = 0.14$ in the regime-switching economy as compared to $k_{med} = 0.1.0311$ in the shutdown economy with the same \bar{H} and $k_{2,0}$). This is because young households have an increased risk in the second period of life under the possibility of a regime switch because their transfer will be lower in the case of a default on \bar{H} .

Table 16 summarizes our findings on time to regime switch; i.e., $w_t \leq H$. Notice that the distributions of time until regime switch across simulations in all the cells of Table 16 are very similar to the distributions in Table 10 where the government takes 80 percent of wages when it can no longer take \bar{H} . Higher precautionary savings extends the time until a regime switch, but increased promised transfers reduce that time.

3.2.2 Fiscal gap and equity premium

Table 17 shows the calculated sure-return prices and their corresponding annualized discount rates for this regime switching economy. Each cell represents the computed prices and interest rates that correspond to a particular promised transfer value \bar{H} and initial capital stock $k_{2,0}$.

		$k_{2,0} = 0.0875$		$k_{2,0} = 0.14$	
		Periods	CDF	Periods	CDF
	\min	1	0.3677	1	0.3340
$\bar{H} = 0.00$	med	2	0.5697	2	0.5440
II = 0.09	mean	3.28	0.7116	3.42	0.7054
	\max	24	1.0000	25	1.0000
	\min	1	0.4517	1	0.4060
$\bar{H} = 0.11$	med	2	0.6390	2	0.6080
II = 0.11	mean	2.80	0.7302	2.96	0.7228
	max	24	1.0000	24	1.0000

Table 16: Periods to regime switch simulationstatistics: 30-percent tax

The "min", "med", "mean", and "max" rows in the "Periods" column represent the minimum, median, mean, and maximum number of periods, respectively, in which the simulated time series hit the regime switch condition. The "CDF" column represents the percent of simulations that switch regimes in t periods or less, where t is the value in the "Periods" column. For the CDF value of the "mean" row, we used linear interpolation.

		$k_{2,0} = 0.0875$		$k_{2,0} =$	= 0.14		
			$r_{t,t+s}$		$r_{t,t+s}$		
	s	$p_{t,t+s}$	APR	$p_{t,t+s}$	APR		
	0	1	0	1	0		
	1	0.3367	0.0370	0.4453	0.0273		
	2	6.0523	-0.0296	8.0476	-0.0342		
$\bar{H} = 0.09$	3	2.0412	-0.0079	6.7823	-0.0210		
	4	8.5075	-0.0177	16.8480	-0.0233		
	5	15.9863	-0.0183	25.3856	-0.0213		
	6	7.5427	-0.0112	6.1479	-0.0100		
	0	1	0	1	0		
	1	0.2326	0.0498	0.3225	0.0384		
	2	7.3132	-0.0326	7.1394	-0.0322		
$\bar{H} = 0.11$	3	11.5166	-0.0268	5.8534	-0.0194		
	4	16.4777	-0.0231	12.1299	-0.0206		
	5	9.2992	-0.0148	15.5375	-0.0181		
	6	23.4145	-0.0174	31.7886	-0.0190		

Table 17: Term structure of prices and in-
terest rates in regime switching
economy: 30-percent tax

The first column in each cell is the price of the sure-return bond $p_{t,t+s}$ at different maturities s as characterized by equation (16). The second column in each cell is the net interest rate $r_{t,t+s}$ APR implied by the sure-return rate and given in annual percentage rate terms according to equation (17). Full descriptions of the term structure of prices and interest rates for all calibrations and for up to s = 12 is provided in the Technical Appendix.

Table 18 shows our four measures of the fiscal gap as a percent of the net present value of GDP for each of our four combinations of \bar{H} and $k_{2,0}$. Similar to the 80percent tax regime switch model, all the measures for the first measure of the fiscal gap (fgap1) are negative. These negative fiscal gaps—and relatively low measures of the fiscal gap for the other measures—occur because the expected H_t after the regime swith is significantly higher than \bar{H} . But in all cases, increased \bar{H} increases the fiscal gap.

Table 18: Measures of the fiscal gap with regime switching as percent of NPV(GDP): 30percent tax

	$k_{2,0} = 0.0875$		$k_{2,0} = 0.14$		
	fgap 1	fgap 2	fgap 1	fgap 2	
	fgap 3	fgap 4	fgap 3	fgap 4	
$\bar{H} = 0.00$	-0.1241	0.0002	-0.1214	-0.0148	
11 = 0.09	0.0099	$\begin{array}{c cccc} 0.0875 & k_{2,0} = 0.14 \\ \hline \text{fgap 2} & \text{fgap 1} & \text{fgap} \\ \hline \text{fgap 4} & \text{fgap 3} & \text{fgap} \\ 0.0002 & -0.1214 & -0.01 \\ 0.0096 & 0.0079 & 0.007 \\ \hline 0.0064 & -0.1190 & -0.01 \\ 0.0171 & 0.0139 & 0.013 \\ \end{array}$	0.0078		
$\bar{H} = 0.11$	-0.1194	0.0064	-0.1190	-0.0108	
II = 0.11	0.0172	0.0171	0.0139	0.0138	

Fiscal gap 1 uses the gross sure return rates $R_{t,t+s}$ from Table 4 as the discount rates for NPV calculation. Fiscal gap 2 uses the current period gross return on capital R_t from the model as the constant discount rate. Fiscal gap 3 uses the International Monetary Fund (2009) method of an annual discount rate equal to 1 plus the average percent change in GDP plus 0.01 (≈ 2.05). And fiscal gap 4 uses the Gohkhale and Smetters (2007) method of an annual discount rate equal to 1 plus 0.0365 (≈ 1.93).

Finally, we calculate the equity premium and Sharpe ratio for this regime switching model. The equity premium results in Table 19 differ little from those in Table 13. This means that the form of the regime change has little effect on the initial period equity premium. The equity premia here around 2 percent with Sharpe ratios around 0.28.

Table 20 reports the average equity premium and Sharpe ratio across simulations in the period immediately before the regime switch as compared to their respective values in the first period. Once again, the average equity premium and Sharpe ratio increase significantly from the initial period to the period right before the regime switch in every case. In both the 80-percent and 30-percent wage-redistribution models, the

		$k_{2,0} = 0.0875$		$k_{2,0} = 0.14$	
		30-year	annual	30-year	annual
	$E[R_{t+1}]$	17.1319	1.0592	12.9708	1.0503
	$\sigma(R_{t+1})$	49.4105	n.a.	37.2570	n.a.
	$R_{t,t+1}$	2.9703	1.0370	2.2457	1.0273
$\bar{H} = 0.09$	Equity premium $E[R_{t+1}] - R_{t,t+1}$	14.1616	0.0223	10.7251	0.0229
	Sharpe ratio $\frac{E[R_{t+1}] - R_{t,t+1}}{\sigma(R_{t+1})}$	0.2866	n.a.	0.2879	n.a.
	$E[R_{t+1}]$	22.1773	1.0678	16.0801	1.0572
	$\sigma(R_{t+1})$	64.1466	n.a.	46.3385	n.a.
	$R_{t,t+1}$	4.2986	1.0498	3.1006	1.0384
$\bar{H} = 0.11$	Equity premium $E[R_{t+1}] - R_{t,t+1}$	17.8787	0.0180	12.9795	0.0187
	Sharpe ratio $\frac{E[R_{t+1}] - R_{t,t+1}}{\sigma(R_{t+1})}$	0.2787	n.a.	0.2801	n.a.

Table 19: Components of the equity premium with
regime switching: 30-percent tax

The gross risky one-period return on capital is $R_{t+1} = 1 + r_{t+1} - \delta$. The annualized gross risky one-period return is $(R_{t+1})^{1/30}$. The expected value and standard deviation of the gross risky one-period return R_{t+1} are calculated as the average and standard deviation, respectively, across simulations. The annual equity premium is the expected value of the annualized risky return in the next period minus the annualized return on the one-period riskless bond. equity premia in the period before the regime switch are close to those observed in the data.

		$k_{2,0} = 0.0875$		$k_{2,0} = 0.14$	
		Eq.	Sharpe	Eq.	Sharpe
		prem.	ratio	prem.	ratio
$\bar{H} = 0.09$	period 1	0.0223	0.2866	0.0229	0.2879
	before shutdown	0.0819	0.3266	0.0848	0.3276
	percent bigger	0.6290	0.5367	0.6617	0.5660
	percent smaller	0.0033	0.0957	0.0043	0.1000
$\bar{H} = 0.11$	period 1	0.0180	0.2787	0.0187	0.2801
	before shutdown	0.0701	0.3173	0.0739	0.3199
	percent bigger	0.5460	0.4807	0.5913	0.5153
	percent smaller	0.0023	0.0677	0.0027	0.0787

Table 20: Equity premium and Sharpe ratio in period immediately before regime switch: 30-percent tax

The "period 1" row represents the equity premium and Sharpe ratio in the initial period for each specification. The "before shutdown" row represents the average equity premium and Sharpe ratio across simulations in the period immediately before shutdown for each specification. The "percent bigger" and "percent smaller" rows tell how many of the simulated ending values of the equity premium and Sharpe ratio were bigger than or less than, respectively, their initial period values. These percentages do not sum to one because the equity premium and Sharpe ratio do not change in the cases in which the economy shuts down in the second period.

4 Conclusion

Our model is as simple as it gets for examining fiscal sustainability. Yet its findings suggest that maintaining unsustainable policies of the kind currently being conducted in the U.S. and other developed nations raises an important set of challenges for long-term economic performance. Younger generations have only 100 percent of their earnings to surrender to older generations. As the government enforces ever greater redistribution, the economy saves and invests less and wages either fall or grow at slower rates than would otherwise be true. In the U.S., generational policy appears responsible for reducing the rate of national saving from roughly 15 percent in the early 1950s to close to zero percent today. The rate of net domestic investment has plunged as well. And for most American workers real wage growth has become a distant memory. Clearly, multi-period models using the sparse grid techniques developed by Krueger and Kubler (2006) are needed to provide more realistic Monte Carlo simulations of actual or near economic death. Whether such models can be developed in time and in sufficient detail to influence developed-country policymakers to alter their current policies remains to be seen.

References

- AGUIAR, M., AND G. GOPINATH (2006): "Defaultable Debt, Interest Rates and the Current Account," *Journal of International Economics*, 69(1), 64–83.
- ALESINA, A., AND A. DRAZEN (1991): "Why are Stabilizations Delayed?," *American Economic Review*, 81(5), 1170–1188.
- ALTIG, D., A. J. AUERBACH, L. J. KOTLIKOFF, K. A. SMETTERS, AND J. WAL-LISER (2001): "Simulating Fundamental Tax Reform in the United States," American Economic Review, 91(3), 574–595.
- ARELLANO, C. (2008): "Default Risk and Income Fluctuations in Emerging Economies," *American Economic Review*, 98(3), 690–712.
- AUERBACH, A. J., AND K. A. HASSETT (1992): "Tax Policy and Business Fixed Investment in the United States," *Journal of Public Economics*, 47(2), 141–170.

(2001): "Uncertainty and the Design of Long-Run Fiscal Policy," in *Demo-graphic Change and Fiscal Policy*, ed. by A. J. Auerbach, and R. D. Lee, pp. 73–92. Cambridge University Press, Cambridge, UK.

(2002): "Fiscal Policy and Uncertainty," International Finance, 5(2), 229–249.

(2007): "Optimal Long-Run Fiscal Policy: Constraints, Preferences, and the Resolution of Uncertainty," *Journal of Economic Dynamics and Control*, 31(5), 1451–1472.

- AUERBACH, A. J., AND L. J. KOTLIKOFF (1987): *Dynamic Fiscal Policy*. Cambridge University Press, Cambridge, MA.
- BARRO, R. J. (2009): "Rare Disasters, Asset Prices, and Welfare Costs," American Economic Review, 99(1), 243–64.
- BLAKE, D. (1996): "Efficiency, Risk Aversion and Portfolio Insurance: An Analysis of Financial Asset Portfolios Held by Investors in the United Kingdom," *Economic Journal*, 106(438), 1175–1192.
- BRAV, A., G. M. CONSTANTINIDES, AND C. C. GECZY (2002): "Asset Pricing with Heterogeneous Consumers and Limited Participation: Empirical Evidence," *Journal of Political Economy*, 110(4), 793–824.
- CAMPBELL, J. Y. (1996): "Understanding Risk and Return," Journal of Political Economy, 104(2), 298–345.

(2000): "Asset Pricing at the Millenium," Journal of Finance, 55(4), 1515-1567.

- COCHRANE, J. H. (2005): Asset Pricing: Revised Edition. Princeton University Press.
- (2011): "Understanding Policy in the Great Recession: Some Unpleasant Fiscal Arithmetic," *European Economic Review*, 55(1), 2–30.
- CONGRESSIONAL BUDGET OFFICE (June 2011): "CBO's Long-Term Budget Outlook," Budget outlook, Congressional Budget Office.
- COOLEY, T. F., AND J. SOARES (1999): "A Positive Theory of Social Security Based on Reputation," *Journal of Political Economy*, 107(1), 135–160.
- DAVIG, T., AND E. M. LEEPER (2011a): "Temporarily Unstable Government Debt and Inflation," NBER Working Paper No. 16799, National Bureau of Economic Research.
- ——— (2011b): "Monetary-Fiscal Policy Interactions and Fiscal Stimulus," *European Economic Review*, 55(2), 211–227.
- DAVIG, T., E. M. LEEPER, AND T. B. WALKER (2010): "Unfunded Liabilities' and Uncertain Fiscal Financing," *Journal of Monetary Economics*, 57(5), 600–619.

(2011): "Inflation and the Fiscal Limit," *European Economic Review*, 55(1), 31–47.

- DE NARDI, M., S. İMROHOROĞLU, AND T. J. SARGENT (1999): "Projected U.S. Demographics and Social Security," *Review of Economic Dynamics*, 2(3), 575–615.
- GOHKHALE, J., AND K. SMETTERS (2007): "Do the Markets Care About the \$2.4 Trillion U.S. Deficit?," *Financial Analysts Journal*, 63(2), 37–47.
- HAMILTON, J. D. (2008): "Regime-switching Models," in New Palgrave Dictionary of Economics, ed. by S. N. Durlauf, and L. E. Blume. Palgrave McMillan Ltd., second edn.
- HASSETT, K. A., AND G. E. METCALF (1999): "Investment with Uncertain Tax Policy: Does Random Tax Policy Discourage Investment?," *The Economic Journal*, 109(457), 372–393.
- HUGGETT, M., AND G. VENTURA (1999): "On the Distributional Effects of Social Security Reform," *Review of Economic Dynamics*, 2(3), 498–531.
- İMROHOROĞLU, A., S. İMROHOROĞLU, AND D. H. JOINES (1995): "A Life Cycle Analysis of Social Security," *Economic Theory*, 6(1), 83–114.

(1999): "Social Security in an Overlapping Generations Model with Land," *Review of Economic Dynamics*, 2(3), 638–665.

- INTERNATIONAL MONETARY FUND (2009): "Fiscal Implications of the Global Economic and Financial Crisis," IMF Staff Position Note SPN/09/13, International Monetary Fund.
- KOCHERLAKOTA, N. R. (1996): "The Equity Premium: It's Still a Puzzle," *Journal* of Economic Literature, 34(1), 42–71.
- KOTLIKOFF, L. J., AND J. GREEN (2009): "On the General Relativity of Fiscal Language," in Key Issues in Public Finance: A Conference in Memory of David Bradford, ed. by A. J. Auerbach, and D. Shaviro, pp. 241–256. Harvard University Press.
- KOTLIKOFF, L. J., K. SMETTERS, AND J. WALLISER (1998a): "The Economic Impact of Transiting to a Privatized Social Security System," in *Redesigning Social Security*, ed. by H. Siebert, pp. 327–348. Kiel University Press, Kiel, Germany.
- (1998b): "Social Security: Privatization and Progressivity," American Economic Review, 88(2), 137–141.
- (2007): "Mitigating America's Demographic Dilemma by Pre-Funding Social Security," *Journal of Monetary Economics*, 54(2), 247–266.
- KRUEGER, D., AND F. KUBLER (2006): "Pareto Improving Social Security Reform when Financial Markets are Incomplete!?," American Economic Review, 96(3), 737–755.
- LEEPER, E. M., AND T. B. WALKER (2011): "Fiscal Limits in Advanced Economies," NBER Working Paper No. 16819, National Bureau of Economic Research.
- MANKIW, N. G., AND S. P. ZELDES (1991): "The Consumption of Stockholders and Nonstockholders," *Journal of Financial Economics*, 29(1), 97–112.
- MEHRA, R., AND E. C. PRESCOTT (1985): "The Equity Premium: A Puzzle," Journal of Monetary Economics, 15(2), 145–161.
- NISHIYAMA, S., AND K. SMETTERS (2007): "Does Social Security Privatization Produce Efficiency Gains?," *Quarterly Journal of Economics*, 122(4), 1677–1719.
- REINHART, C. M., AND K. ROGOFF (2009): This Time is Different: Eight Centuries of Financial Folly. Princeton University Press.
- SARGENT, T. J., AND N. WALLACE (1981): "Some Unpleasant Monetarist Arithmetic," Federal Reserve Bank of Minneapolis Quarterly Review, 5(3), 1–17.
- SHILLER, R. J. (1982): "Consumption, Asset Markets, and Macroeconomic Fluctuations," Carnegie-Rochester Conference Series on Public Policy, 17(1), 203–238.
- SMETTERS, K., AND J. WALLISER (2004): "Opting Out of Social Security," Journal of Public Economics, 88(7-8), 1295–1306.

- TRABANDT, M., AND H. UHLIG (2009): "How Far Are We from the Slippery Slope? The Laffer Curve Revisited," NBER Working Paper No. 15343, National Bureau of Economic Research.
- YUE, V. Z. (2010): "Sovereign Default and Debt Renegotiation," Journal of International Economics, 80(2), 176–187.

APPENDIX

A-1 Description of calibration

This section details how we arrived at the calibrated parameter values listed in Table 1. The 30-year discount factor β is set to match the annual discount factor common in the RBC literature of 0.96.

$$\beta = (0.96)^{30}$$

We set the coefficient of relative risk aversion at a midrange value of $\gamma = 2$. This value lies in the midrange of values that have been used in the literature.¹² The capital share of income parameter is set to match the U.S. average $\alpha = 0.35$, and the 30-year depreciation rate δ is set to match an annual depreciation rate of 5 percent.

$$\delta = 1 - (1 - 0.05)^{30}$$

The equilibrium production process in our 2-period model is the following,

$$Y_t = e^{z_t} K_t^{\alpha} \quad \forall t$$

where labor is supplied inelastically and z_t is the aggregate total factor productivity shock. We assume the shock z_t is an AR(1) process with normally distributed errors.

$$z_t = \rho z_{t-1} + (1-\rho)\mu + \varepsilon_t$$

where $\rho \in [0,1), \quad \mu \ge 0, \quad \text{and} \quad \varepsilon_t \sim N(0,\sigma^2)$ (5)

This implies that the shock process e^{z_t} is lognormally distributed $LN(0, \sigma^2)$. The RBC literature calibrates the parameters on the shock process (5) to $\rho = 0.95$ and $\sigma = 0.4946$ for annual data.

For data in which one period is 30 years, we have to recalculate the analogous $\tilde{\rho}$ and $\tilde{\sigma}$.

$$\begin{aligned} z_{t+1} &= \rho z_t + (1-\rho)\mu + \varepsilon_{t+1} \\ z_{t+2} &= \rho z_{t+1} + (1-\rho)\mu + \varepsilon_{t+2} \\ &= \rho^2 z_t + \rho(1-\rho)\mu + \rho\varepsilon_{t+1} + (1-\rho)\mu + \varepsilon_{t+2} \\ z_{t+3} &= \rho z_{t+2} + (1-\rho)\mu + \varepsilon_{t+3} \\ &= \rho^3 z_t + \rho^2(1-\rho)\mu + \rho^2\varepsilon_{t+1} + \rho(1-\rho)\mu + \rho\varepsilon_{t+2} + (1-\rho)\mu + \varepsilon_{t+3} \\ &\vdots \\ z_{t+j} &= \rho^j z_t + (1-\rho)\mu \sum_{s=1}^j \rho^{j-s} + \sum_{s=1}^j \rho^{j-s}\varepsilon_{t+s} \end{aligned}$$

¹²Estimates of the coefficient of relative risk aversion γ mostly lie between 1 and 10. See Mankiw and Zeldes (1991), Blake (1996), Campbell (1996), Kocherlakota (1996), Brav, Constantinides, and Geczy (2002), and Mehra and Prescott (1985).

With one period equal to thirty years j = 30, the shock process in our paper should be:

$$z_{t+30} = \rho^{30} z_t + (1-\rho)\mu \sum_{s=1}^{30} \rho^{30-s} + \sum_{s=1}^{30} \rho^{30-s} \varepsilon_{t+s}$$
(A.1.1)

Then the persistence parameters in our one-period-equals-thirty-years model should be $\tilde{\rho} = \rho^{30} = 0.2146$. Define $\tilde{\varepsilon}_{t+30} \equiv \sum_{s=1}^{30} \rho^{30-s} \varepsilon_{t+s}$ as the summation term on the right-hand-side of (A.1.1). Then $\tilde{\varepsilon}_{t+30}$ is distributed:

$$\tilde{\varepsilon}_{t+30} \sim N\left(0, \left[\sum_{s=1}^{30} \rho^{2(30-s)}\right]\sigma^2\right)$$

Using this formula, the annual persistence parameter $\rho = 0.95$, and the annual standard deviation parameter $\sigma = 0.4946$, the implied thirty-year standard deviation is $\tilde{\sigma} = 1.5471$. So our shock process should be,

$$z_t = \tilde{\rho} z_{t-1} + (1-\rho)\tilde{\mu} + \tilde{\varepsilon}_t \quad \forall t \quad \text{where} \quad \tilde{\varepsilon} \sim N(0, \tilde{\sigma}^2)$$

where $\tilde{\rho} = 0.2146$ and $\tilde{\sigma} = 1.5471$. We calibrate μ , and therefore $\tilde{\mu}$, so that the median wage is 50,000.

Lastly, we set the size of the promised transfer \overline{H} to be 32 percent of the median real wage. This level of transfers is meant to approximately match the average per capita real transfers in the United States to the average real wage in recent years. We get the median real wage by simulating a time series of the economy until it hits the shut down point, and we do this for 3,000 simulated time series. We take the median wage from those simulations. In order to reduce the effect of the initial values on the median, we take the simulation that lasted the longest number of periods before shutting down and remove the first 10 percent of the longest simulation's periods from each simulation for the calculation of the median.